

## 14 Inference in complex models

What if graph is not a tree?

NP-hard even to approximate marginals and conditionals

### General strategies

1. Exact methods — exponential time, but can still try to be smart
2. Approximation methods
3. Heuristic methods
4. Monte Carlo methods — estimate by random sampling

### 14.1 Exact methods

#### Elimination ordering

Try to find a good variable order that reduces work in summation

- push variable in
- eliminate variables by summing and pull result out

#### Variable clustering

Cluster variables to create a tree structured Bayesian network

- exponential in the size of the largest cluster

#### Cut sets

Choose a cut set of variables that turn factor graph into a tree

- sum over cut set configurations
- exponential in size of cut set

## 14.2 Approximation methods

“Variational approximation”

- Pick simple model structure (i.e. a tree)
- Set values in new CP tables so that new distribution approximates original distribution as closely as possible
- Perform efficient inference on simpler approximate distribution

A bit complicated to implement sometimes, but can be very effective

## 14.3 Heuristic methods

“Loopy probability propagation”

Ignore loops and use same message passing algorithm as for trees

- random initial messages
- keep passing messages around graph
- wait for product of incoming messages to converge
- if so, is the answer accurate?

This works way better than it should!

## 14.4 Monte Carlo methods

Use random sampling to *estimate* answers

### 14.4.1 Estimating marginals

To estimate  $P(X_i = x_i)$ , draw joint configurations

$$\begin{array}{cccc} x_{11} & x_{12} & \dots & x_{1n} \\ x_{21} & x_{22} & \dots & x_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ x_{t1} & x_{t2} & \dots & x_{tn} \end{array}$$

Use estimate:  $\hat{P}(X_i = x_i) = \frac{\# \text{ matches}(X_i = x_i)}{t}$

Unbiased:  $E\hat{P}(X_i = x_i) = P(X_i = x_i)$

#### 14.4.2 Estimating conditionals

Estimate  $P(X_{k+1} = y_{k+1} | X_1 = x_1, \dots, X_k = x_k)$

Draw joint configurations:

$$\begin{array}{cccc} x_{11} & x_{12} & \dots & x_{1n} \\ x_{21} & x_{22} & \dots & x_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ x_{t1} & x_{t2} & \dots & x_{tn} \end{array}$$

Use estimate:

$$\begin{aligned} \hat{P}(X_{k+1} = y_{k+1} | X_1 = x_1, \dots, X_k = x_k) \\ = \frac{\# \text{ matches}(X_1 = x_1, \dots, X_k = x_k, X_{k+1} = y_{k+1})}{\# \text{ matches}(X_1 = x_1, \dots, X_k = x_k)} \end{aligned}$$

This technique is called “logic sampling”

It is a bad estimator if  $(X_1 = x_1, \dots, X_k = x_k, X_{k+1} = y_{k+1})$  is unlikely:

- small effective sample size

#### 14.4.3 Aside: General “importance sampling”

Consider estimating the expected value of some function  $f(x)$ , where  $x$  is drawn randomly according to the distribution  $P(x)$ . That is, assume the expectation of  $f(x)$  is defined

$$E_{P(x)}(f(x)) = \sum_x f(x)P(x)$$

Many problems (including estimating conditional probabilities) can be expressed as estimating the expected value of a function  $f$ .

The simplest way to estimate  $E_{P(x)}f(x)$  is the Monte Carlo method

- Draw  $x_1, x_2, \dots, x_t$  from  $P$
- Use estimate:

$$\hat{f} = \frac{1}{t} \sum_{i=1}^t f(x_i)$$

Problem: what if you cannot sample from  $P$  efficiently?

First assume that we can at least efficiently *evaluate*  $P(x)$  at given points  $x$ .

**Idea:** Pick a proposed distribution  $Q$  that you *can* sample from

- Draw  $x_1, x_2, \dots, x_t$  from  $Q$ .
- Weight points by  $w(x_i) = \frac{P(x_i)}{Q(x_i)}$
- Use estimate:  $\hat{f} = \frac{1}{t} \sum_{i=1}^t f(x_i)w(x_i)$

This gives an unbiased estimate

$$\begin{aligned}
 \frac{1}{t} \sum_{i=1}^t f(x_i)w(x_i) &\xrightarrow{t \rightarrow \infty} E_{Q(x)} f(x)w(x) \\
 &= \sum_x f(x)w(x)Q(x) \\
 &= \sum_x f(x) \frac{P(x)}{Q(x)} Q(x) \\
 &= \sum_x f(x)P(x) \\
 &= E_{P(x)} f(x).
 \end{aligned}$$

**More realistically:** You cannot even *evaluate*  $P(x)$  efficiently

However, in these cases, you often still have a function  $R(x) = \beta P(x)$  that you can evaluate efficiently (up to some unknown value  $\beta$ ). In which case you can use following *indirect* importance sampling procedure.

- Draw  $x_1, x_2, \dots, x_t$  from  $Q$ .
- Weight points by  $u(x) = \frac{R(x)}{Q(x)}$
- Use the estimate

$$\hat{f} = \frac{\sum_{i=1}^t f(x_i)u(x_i)}{\sum_{i=1}^t u(x_i)}$$

This procedure is biased, but it is asymptotically unbiased:

$$\frac{1}{t} \sum_{i=1}^t f(x_i)u(x_i) \xrightarrow{t \rightarrow \infty} \sum_x f(x)u(x)Q(x) = \sum_x f(x)R(x) = \beta \sum_x f(x)P(x)$$

$$\frac{1}{t} \sum_{i=1}^t u(x_i) \xrightarrow{t \rightarrow \infty} \sum_x u(x) Q(x) = \sum_x R(x) = \beta \sum_x P(x) = \beta$$

Therefore

$$\hat{f} \xrightarrow{t \rightarrow \infty} \frac{\beta \sum_x f(x) P(x)}{\beta} = E_{P(x)} f(x).$$

#### 14.4.4 Estimating conditionals using importance sampling

Want to estimate  $P(\mathbf{x}_\beta = \mathbf{y}_\beta | \mathbf{x}_\alpha = \mathbf{x}_\alpha)$  where  $\alpha$  and  $\beta$  are sets of indices from  $\{1, \dots, n\}$  such that  $\alpha \cap \beta = \emptyset$  and  $\alpha \cup \beta = \{1, \dots, n\}$ . unfortunately it is both hard to sample from and evaluate  $P(\mathbf{x}_\beta = \mathbf{y}_\beta | \mathbf{x}_\alpha = \mathbf{x}_\alpha)$  directly. we proceed as follows

- clamp the variables  $\mathbf{x}_\alpha = \mathbf{x}_\alpha$
- sample the remaining “free” variables in the usual way (keeping the clamped variables at their assigned values)
- repeat  $t$  times to create a sample of configurations  $\mathbf{x}_1, \dots, \mathbf{x}_t$
- Define the function

$$f(\mathbf{x}_\beta) = \begin{cases} 1 & \text{if } \mathbf{x}_\beta = \mathbf{y}_\beta \\ 0 & \text{otherwise} \end{cases}$$

- Calculate weights

$$u(\mathbf{x}_{\beta,i}) = \frac{R(\mathbf{x}_{\beta,i})}{Q(\mathbf{x}_{\beta,i})}$$

where  $R(\mathbf{x}_{\beta,i}) = P(\mathbf{X}_\alpha = \mathbf{x}_{\alpha,i}, \mathbf{X}_\beta = \mathbf{x}_{\beta,i})$

and  $Q(\mathbf{x}_{\beta,i}) = \prod_{j \in \beta} P(X_j = x_{j,i} | \mathbf{X}_{\pi(j)} = \mathbf{x}_{\pi(j),i})$

- Use the estimate

$$\hat{P}(\mathbf{x}_\beta = \mathbf{y}_\beta | \mathbf{x}_\alpha = \mathbf{x}_\alpha) = \frac{\sum_{i=1}^t f(\mathbf{x}_{\beta,i}) u(\mathbf{x}_{\beta,i})}{\sum_{i=1}^t u(\mathbf{x}_{\beta,i})}$$

This method has larger effective sample size than logic sampling.

Works even if  $P(\mathbf{X}_\alpha = \mathbf{x}_\alpha)$  is small.

## **Readings**

Russell and Norvig: Section 14.5

Dean, Allen, Aloimonos: Section 8.3